

SANS2022-23-HW1

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1 Multivariate Gaussian random variables

Generate for $N = 100, 1000, 10000$ samples of a multivariate gaussian variable $X = (X_1, X_2)$ with $\mu = (1, 2)$, $\Sigma = \text{diag}(1, 4)$. Plot the scatterplots of the generated samples for the three values of N .

Plot the histograms for X_1 and X_2 . Can you recognize distribution of the corresponding random variable? Justify your answer taking into account the theory on gaussian random variables.

Repeat for Σ given by:

$$\Sigma = \begin{bmatrix} \frac{\sqrt{2}}{2} & -\frac{\sqrt{2}}{2} \\ \frac{\sqrt{2}}{2} & \frac{\sqrt{2}}{2} \end{bmatrix} \begin{bmatrix} 1 & 0 \\ 0 & 4 \end{bmatrix} \begin{bmatrix} \frac{\sqrt{2}}{2} & \frac{\sqrt{2}}{2} \\ -\frac{\sqrt{2}}{2} & \frac{\sqrt{2}}{2} \end{bmatrix} \quad (1)$$

Explain what you see

2 Bayesian estimation of 2D Gaussian random variables

Note: For solving this, you need to find somewhere the adequate formulas. In the class we only explained the 1D case, and very briefly the multidimensional case. You are in a Master now.

We want to estimate the location of an static point located at point $(2, 3)$. We measure $N = 10$ samples of the location of the point. The measure error has a known variance-covariance matrix given by:

$$\Sigma = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \quad (2)$$

Generate $N = 10$ samples of our reading. Draw the scatter plot.

Estimate μ . Use a uniform prior for μ . Draw 3-D plot of the likelihood and posterior distributions.

Repeat, but assuming now a gaussian prior for μ with different values of the mean μ_0 and variance-covariance matrix Σ_0 . Draw also 3-D plots for these priors.

3 Metropolis-Hastings Monte Carlo simulation of a multinomial distribution in a simplex

Note: you can use your own code or use the code I've wrote and that you can find in the raco. I'm not a great python programmer, sorry .

We want to estimate the posterior distribution $p(\theta|y)$ and predictive probability $p(y_{n+1} = 1|y)$ of a multinomial distribution. In this case there are closed expressions for these distributions, but we want to use MCMC.

For the auxiliary Markov-process that generates candidate next states, we start using a uniform distribution: $Q(\theta, \theta') = \text{constant}$ for all values of θ' in the simplex, and $Q(\theta, \theta') = 0$ for θ' outside the simplex. For generating samples from such uniform distribution we use *Dirichlet*((1, 1, 1, ..., 1)).

Run for $T = 100, 1000, 10000$ the corresponding M-H simulation for different values of y , for instance:

$$y = [2/7, 1/7, 4/7],$$

$$y = [2/15, 1/15, 1/15, 1/15, 1/15, 2/15, 1/15, 0, 2/15, 1/15, 3/15].$$

Plot the obtained histograms for the averages of the marginals of the posterior $\mathbb{E}(\theta_i|y)$ and the predictive probability $p(y_{n+1} = 0|y)$. Include also in the plots the true values of the estimated parameter.

For computing η , the acceptance probability $\eta = \min(1, \frac{p^*(\theta'|y)}{p^*(\theta|y)})$, in the code I use

$$\eta = \min(1, \exp(\log(p^*(\theta'|y)) - \log(p^*(\theta|y)))),$$

why?.

This choice of Q has a clear problem: it generates too many candidates that are not taken (i.e. the rejection rate in the algorithm is too high). Explain why.

Think in another distribution for Q . In the code in the raco you can find a distribution I've tried. It seems to work, explain why. What is the meaning of the parameter h ?. Try with $h = 1, 0.1, 0.001$ and explain the results.